

# Advanced statistical methods and Bayesian inference in scientific research

## Lecture 4

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## Random variable transformation

Let transform  $X$  to  $Y$

$$X \implies Y$$

Consistency requirement:

$$\forall_{\mathcal{A}} \ P_X(\mathcal{A}) = P_Y(\mathcal{B})$$

$$p(y) = p(x) \left| \frac{\partial x}{\partial y} \right|$$

## Bayes theorem - single variable X

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Bayes “theorem”

$$P(B|A) = P(A|B) \times P(B)/P(A)$$

## Bayes theorem - application

- ◆ random variable  $X$  is described by  $\rho(x)$
- ◆ we perform another measurement of  $X$

Question: how performed measurement allows us to constraint (update) the pdf distribution  $p(x)$  describing X?

## Bayes theorem - single variable X

$$P(B|A) = P(A|B) \times P(B)/P(A)$$

Interpretation:

A → measured values of X

B → values of X drawn from  $\rho(x)$

$P(A)$  → “normalization term - see next slide”

$P(B)$  → “initial” pdf  $\rho(x)$

$P(B|A)$  → thought  $p(x)$

$P(A|B)$  → how possible values of  $x$  relates to measured  $x_o$  expressed by  $l(x_o|x)$

$$p^{(1)}(x) = \kappa l(x_o|x) \times \rho(x)$$

when more measurements becomes available

$$p^{(n)}(x) = \kappa l(x_o, x_1, \dots, x_{n-1}|x) \times p^{(n-1)}(x)$$

## Bayes theorem - single variable X

Normalization condition (symbolic):

$$\int P(B|A)dB = 1$$

$$P(A) = \int P(A|B) \times P(B)$$

Problematic contradiction in interpretation of  $P(A)$  and  $P(B)$

## Mathematics of inference - Inference Space

$$(\mathcal{P}, \Sigma, \wedge)$$

where

$\mathcal{P}$  - parameter space

$\Sigma$  - space of all probability distributions over  $\mathcal{P}$

$\wedge(\cdot, \cdot)$  - joining operator:  $\Sigma \times \Sigma \rightarrow \Sigma$

## Joining information according to Tarantola

Two distributions describing **different** pieces of information about the same object

1.  $p(x)$
2.  $q(x)$

$$\zeta(x) = \mathbf{p} \wedge \mathbf{q}(\mathbf{x}) = \frac{\mathbf{p}(\mathbf{x}) \mathbf{q}(\mathbf{x})}{\mu(\mathbf{x})}$$

$\mu(x)$  - non-informative probability

## Non-informative distribution

$$q(x) = \mu(x)$$

$$\mathbf{p} \wedge \mathbf{q}(\mathbf{x}) = \frac{\mathbf{p}(\mathbf{x}) \mu(\mathbf{x})}{\mu(\mathbf{x})} = \mathbf{p}(\mathbf{x})$$

Essentially,  $\mu(\cdot)$  can be arbitrary but usually is taken as  
*volumetric pdf*

## Using Inference space approach - example I

- ◆ random variable  $X$  is described by  $\rho(x)$
- ◆ we perform another measurement of  $X$

Question: how performed measurement allows us to constraint (update) the pdf distribution  $p(x)$  describing  $X$ ?

Assumption: noisy measurement with known noise characteristic  $\psi(\cdot)$

$$x_o = x_o^{true} + n$$

$$X_o \implies p(x) = \psi(x - x_o)$$

## Using Inference space approach - example I

$$p \wedge q(x) = \frac{p(x) \mu(x)}{\mu(x)} = p(x)$$

$$p^{(1)}(x) = \frac{1}{Z} \frac{\rho(x)\psi(x - x_o)}{\mu(x)}$$

$$Z = \int \rho(x)\psi(x - x_o)/\mu(x)dx$$

Comment:

$Z$  (evidence) measure to what extend measurement is compatible with “apriori”  $\rho(x)$

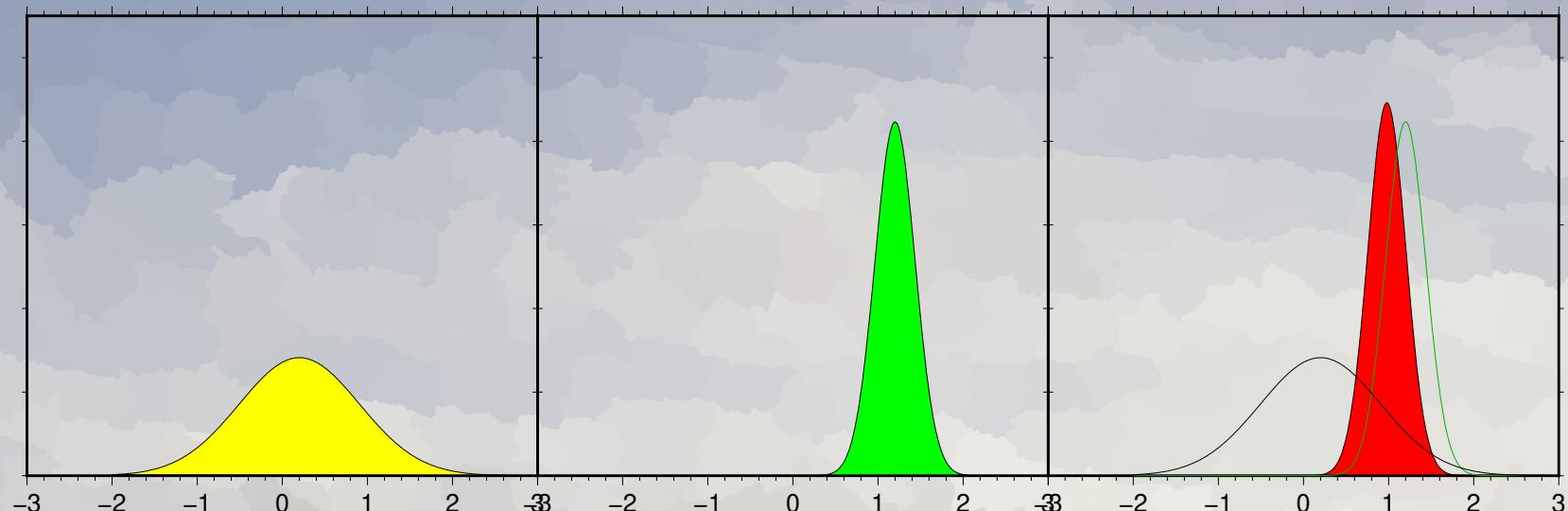
## Using Inference space approach - example I

If another (independent) measurement is available

$$p^{(2)}(x) = \frac{1}{Z'} \frac{p^{(1)}(x)\psi(x - x_o)}{\mu(x)}$$

$$p^{(n)}(x) = \frac{1}{Z} \frac{p^{(n-1)}(x)\psi(x - x_o^{(n)})}{\mu(x)}$$

# Using Inference space approach - example I



## Using Inference space approach - example II

- ◆ random variable  $X$  is described by  $\rho(x)$
- ◆ we perform a measurement of  $Y$
- ◆ measurement errors are characterized by  $\psi(y - y_o)$  distribution
- ◆ we know that  $X$  and  $Y$  are related as  $Y = G(X) + \epsilon$   
and the relation is subjected to errors  $\epsilon$  described by

$$\zeta(X, Y) = \zeta(Y - G(X))$$

Question:

how performed measurement constraints (update) knowledge of  $X$  ?

## Using Inference space approach - example II

New “vectorized” random variable

$$X, Y \implies Z := \begin{pmatrix} X \\ Y \end{pmatrix}$$

$$\mathcal{R} \times \mathcal{R} \implies \mathcal{R}^2$$

$$\rho(x) \rightarrow \rho'(z) = \rho(x)\mu(y)$$

$$\psi(y - y_o) \rightarrow \psi'(z) = \psi(y - y_o)\mu(x)$$

$$\zeta(x, y) \rightarrow \zeta'(z)$$

## Using Inference space approach - example II

$$p \wedge q(x) = \frac{p(x) q(x)}{\mu(x)}$$

$$p(z) = a \frac{\psi'(z)\zeta'(z)}{\mu(z)}$$

$$\sigma(z) = a \frac{p(z)\rho'(z)}{\mu(z)}$$

$$\sigma(z) = \frac{1}{Z} \frac{\zeta(z) \psi'(z) \rho'(z)}{\mu^2(z)}$$

## Using Inference space approach - example II

Taking marginal integrals

$$\sigma(x) = \int_Y \sigma(z) dy$$

$$\sigma(x) = \frac{1}{Z} \rho(x) \int_Y \psi(y - y_o) \zeta(y - G(x)) dy$$

$$Z = \int_X \int_Y \rho(x) \psi(y - y_o) \zeta(y - G(x)) dy dx$$

## Using Inference space approach - example II

If the relation between  $X$  and  $Y$  is exactly known

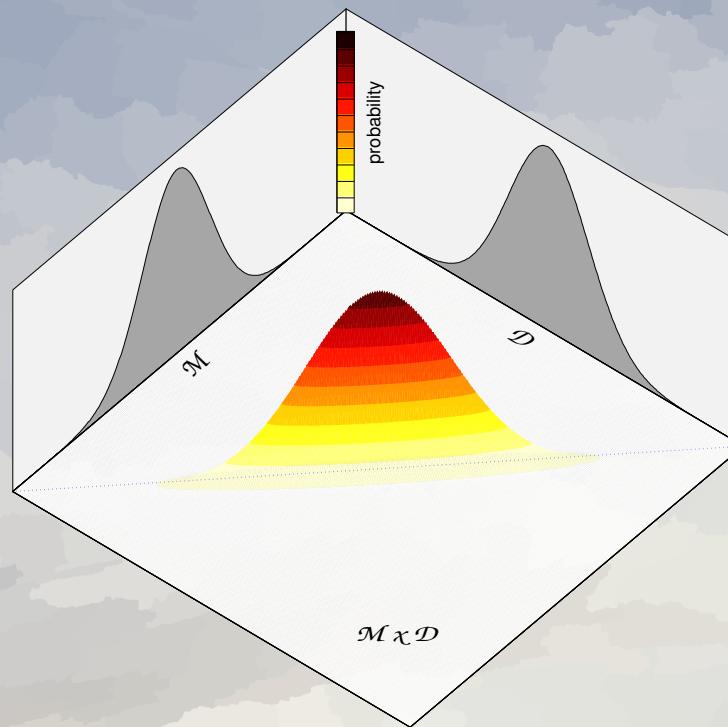
$$\zeta(y - G(x)) = \delta(y - G(x))$$

$$\sigma(x) = \frac{1}{Z} \rho(x) \psi(y_o - G(x))$$

$$p^{(1)}(x) = \frac{1}{Z} \frac{\rho(x)\psi(x_o - x)}{\mu(x)}$$

## Using Inference space approach - example II

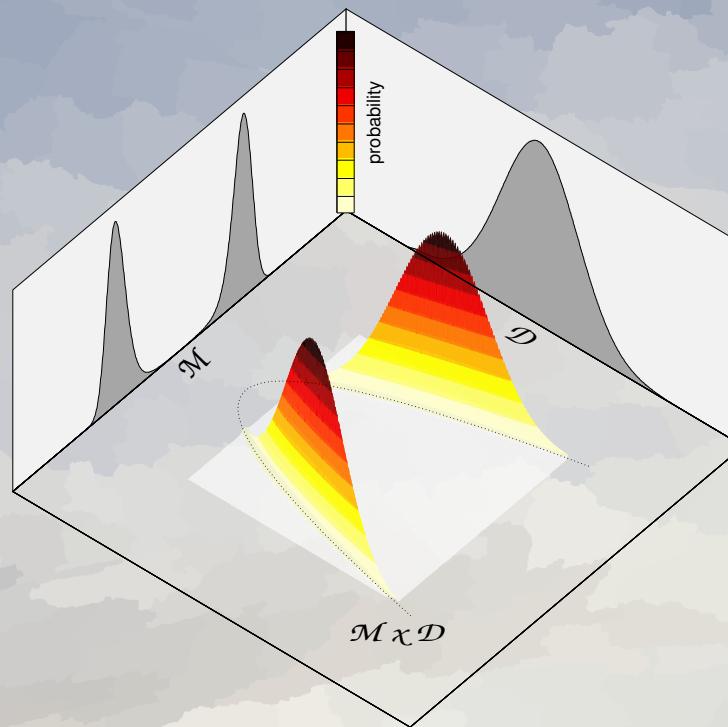
Exact linear relation  $y = G \cdot x$



$$\sigma(x) = \frac{1}{Z} \rho(x) \psi(y_o - G \cdot x)$$

## Using Inference space approach - example II

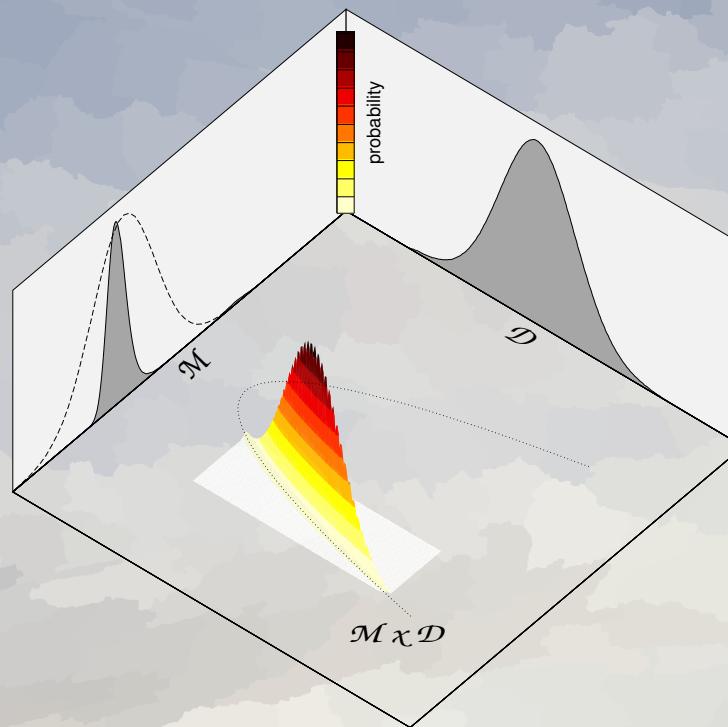
Exact non-linear relation  $y = G \cdot x^2$



$$\sigma(x) = \frac{1}{Z} \rho(x) \psi(y_o - G \cdot x^2)$$

## Using Inference space approach - example II

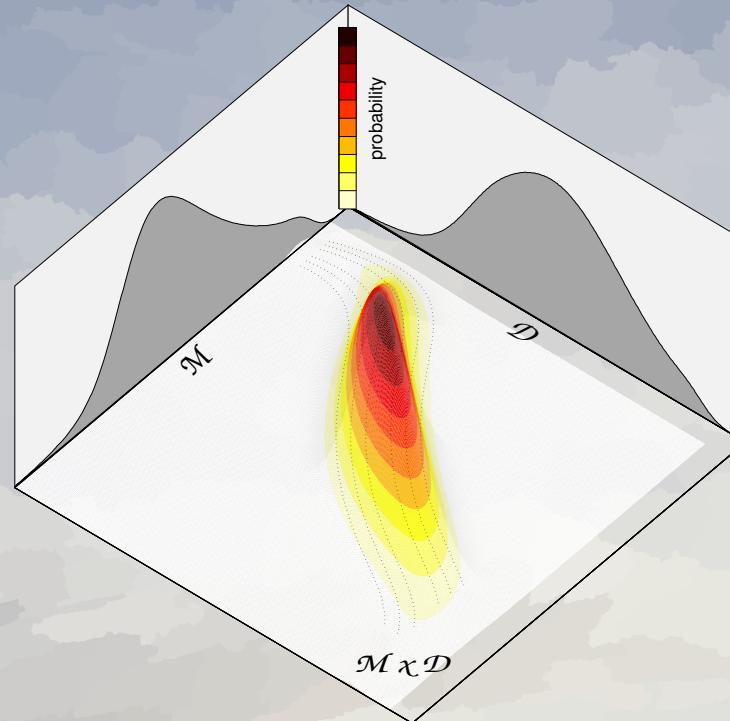
Exact non-linear relation  $y = G \cdot x^2$  and “*a priori*” term



$$\sigma(x) = \frac{1}{Z} \rho(x) \psi(y_o - G \cdot x^2)$$

# Using Inference space approach - example II

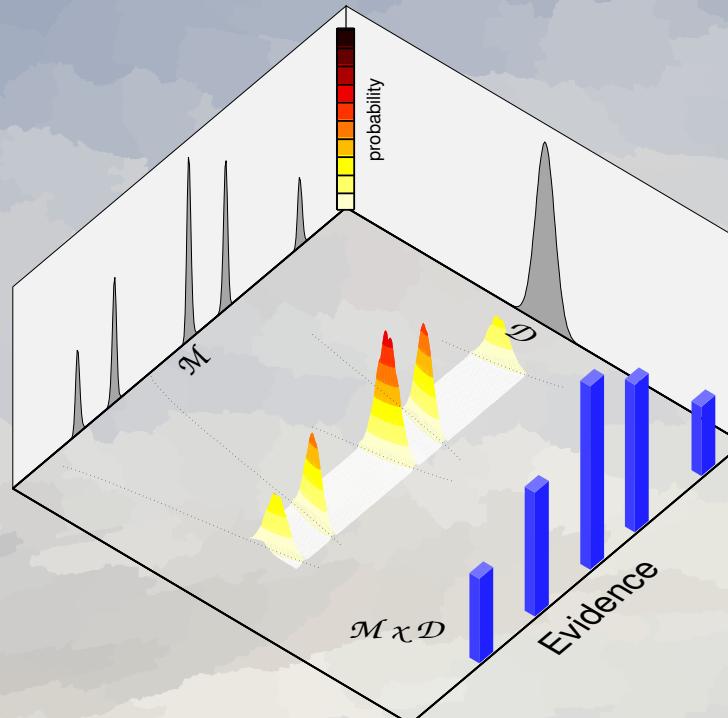
## Approximate theory



$$\sigma(x) = \frac{1}{Z} \rho(x) \int_Y \psi(y - y_o) \zeta(y - G(x)) dy$$

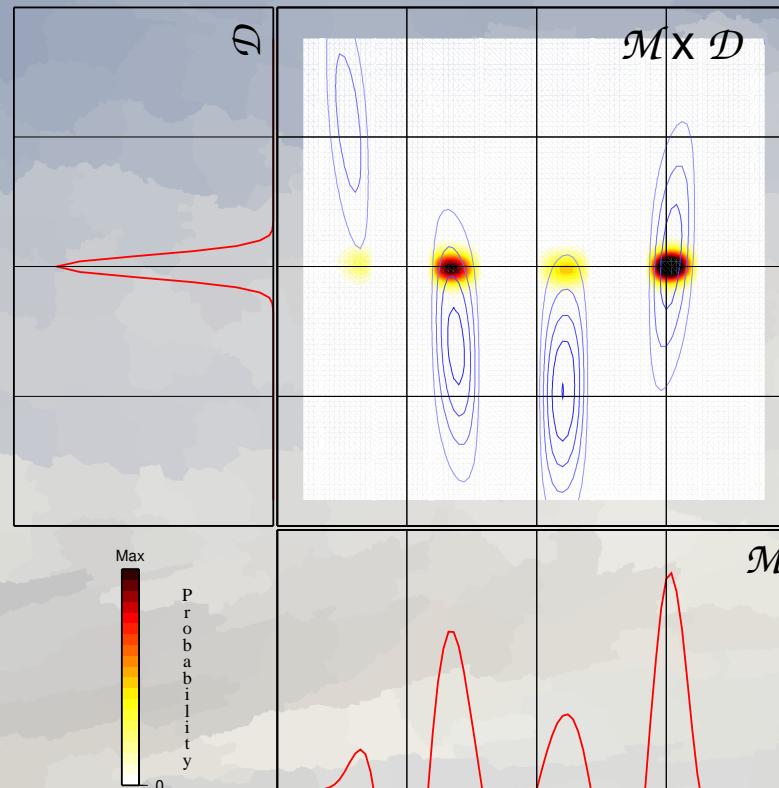
# Using Inference space approach - example II

## Comparing various theories



# Using Inference space approach - example II

## Comparing various theories



# Evidence

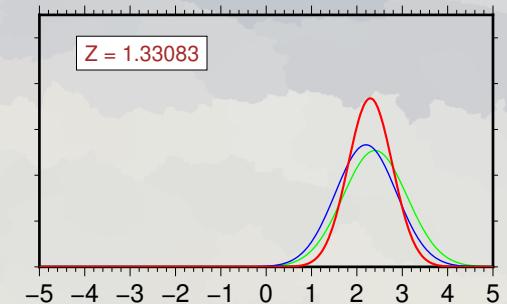
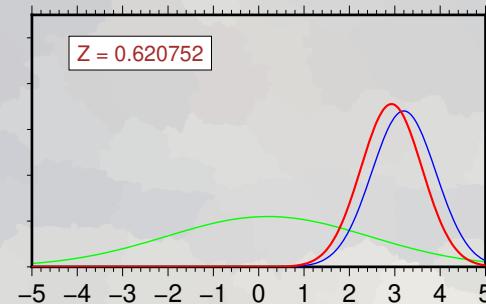
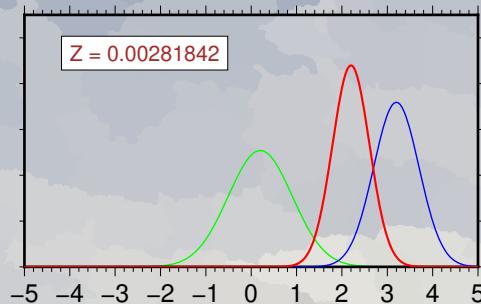
$$Z = \int_X \int_Y \rho(x) \psi(y - y_o) \zeta(y - G(x)) dy dx$$

$$\mathcal{L}(x) = \int_Y \psi(y - y_o) \zeta(y - G(x)) dy$$

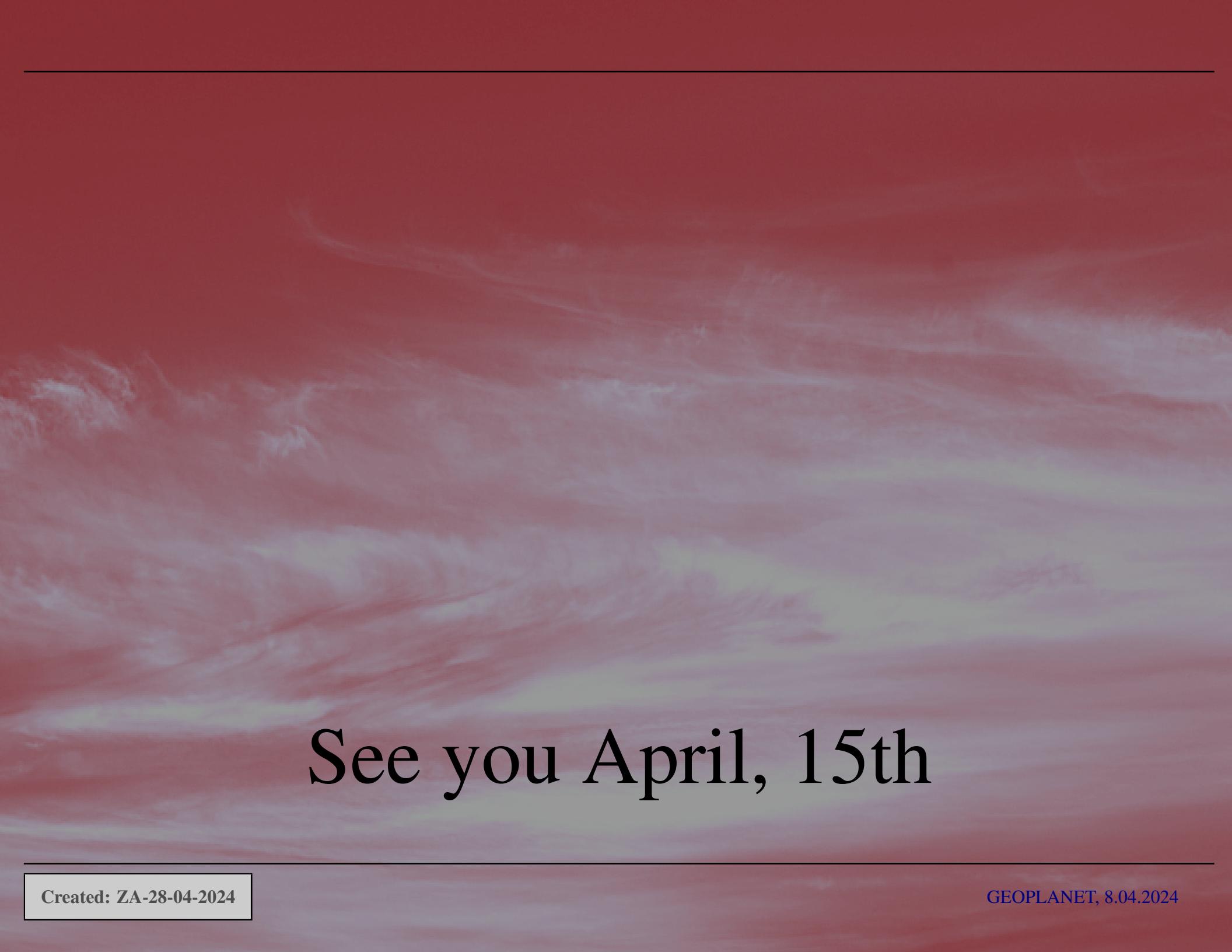
$$Z = \int_X \rho(x) \mathcal{L}(x) dx$$

# Evidence

- ◆  $\rho(x)$  - green
- ◆  $\mathcal{L}(x)$  - blue
- ◆  $\sigma(x) = \rho(x)\mathcal{L}(x)$  - red



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See you April, 15th